# **A comprehensive framework for the study of species co-occurrences, nestedness and turnover**

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Binary presence–absence matrices (rows  $=$  species, columns  $=$  sites) are often used to quantify patterns of species co-occurrence, and to infer possible biotic interactions from these patterns. Previous classifccations of co-occurrence patterns as nested, segregated, or modular have led to contradictory results and conclusions. ese analyses usually do not incorporate the functional traits of the species or the environmental characteristics of the sites, even though the outcomes of species interactions often depend on trait expression and site quality. Here we address this shortcoming by developing a method that incorporates realized functional and environmental niches, and relates them to species co-occurrence patterns. ese niches are defined from n-dimensional ellipsoids, and calculated from the n eigenvectors and eigenvalues of the variance– covariance matrix of measured environmental or trait variables. Average niche overlap among species and the spatial distribution of niches define a triangle plot with vertices of species segregation (low niche overlap), nestedness (high niche overlap), and modular co-occurrence (clusters of overlapping niches). Applying this framework to temperate understorey plant communities in southwest Poland, we found a consistent modular structure of species occurrences, a pattern not detected by conventional presence–absence analysis. ese results suggest that, in our case study, habitat filtering is the most important process structuring understorey plant communities. Furthermore, they demonstrate how incorporating trait and environmental data into co-occurrence analysis improves pattern detection and provides a stronger theoretical framework for understanding community structure.

Since the pioneering work of Diamond (1975), the assembly of species communities has remained a central research focus (Weiher and Keddy 1999, Stokstad 2009, Götzenberger et al. 2012).  $\qquad$  e core data structure – a species  $\times$  sites binary presence–absence matrix (Connor and Simberloff 1979) – has been used to infer species interactions (Fari a et al. 2009), the distribution of species functional traits (Kraft et al. 2008), and stochastic colonisation and extinction (Hubbell 2001).

Diamond (1975) and subsequent authors (reviewed by Weiher and Keddy 1999, Chesson, 2000, Götzenberger et al. 2012) focused on non-random patterns of compositional change across replicate communities. ev interpreted a pattern of reciprocal species occurrence (species segregation) as evidence for competitive exclusion (Gause 1934). But, unless habitats are identical in the replicate patches and there are no limits to dispersal, a pattern of species segregation cannot be distinguished from a change in species composition driven by a gradient in habitat quality or by limited dispersal (species turnover; Ulrich and Gotelli 2013, Blois et al. 2014). Subsequent work (reviewed by

Ulrich and Gotelli 2013) showed that a pattern of reciprocal exclusion in a presence–absence matrix is statistically indistinguishable from a pattern of regular change in species composition (species turnover). us, species segrecies composition (species turnover). gation and turnover appeared to be two sides of the same coin. Any highly segregated matrix can be reordered to show its intrinsic high degree of species turnover among sites (Fig. 1). Recent meta-analyses of animal and plant communities consistently reported a tendency towards species segregation among modern assemblages sampled at a variety of spatial scales (Gotelli and McCabe 2002, Ulrich and Gotelli 2007, 2013), but not among pre-Holocene assemblages (Lyons et al. 2016).

An alternative, biogeographic concept of community assembly focuses on nested patterns of species co-occurrences (Hultén 1937), in which species-poor communities are perfect nested subsets of species-rich communities (Patterson and Atmar 1986, reviewed in Ulrich et al. 2009). A nested community pattern (Fig. 1) is expected from stochastic colonization and extinction in the equilibrium model of island biogeography (MacArthur and Wilson 1963) as well

as from stochastic speciation, dispersal, and extinction in ecological drift models (Hubbell 2001). Furthermore, nested patterns might also emerge in the presence of ecological gradients promoting the co-existence of habitat generalists and habitat specialists. For this reason, Ulrich et al. (2009) recommended nestedness analysis as a major tool to identify environmental gradients along which local communities are assembled in an orderly manner.

Species turnover and nestedness have been interpreted as the endpoints of a continuum along which meta-communities are organised (Presley et al. 2010, Brown et al. 2017). Here, we define a meta-community as a set of local patches that are connected by dispersal, and in which all species have the potential to occur (Wilson 1992). However, it is still unclear whether meta-communities really can be ordered according to this assumed one-dimensional continuum from turnover to nestedness. For example, Ulrich and Gotelli (2013) illustrated how presence–absence matrices identified by standard metrics can be simultaneously nested and segregated. Similar results were also obtained in analyses of trophic networks: Fortuna et al. (2010) reported both positive and negative correlations of metrics quantifying nestedness and segregation. Moreover, the structure of these correlations changed with the fill of the matrix.

e natural counterpart of segregation is modularity, a pattern in which groups of species frequently occur in the between-module turnover. In contrast, a nested pattern is a single module in which species occurrence frequencies vary, and composition is nested.

# **Environmentally defined patterns of co-occurrence**

e classical approach of analysing species co-occurrence with only a binary presence–absence matrix has three important drawbacks. First, it does not consider environmental constraints (Leibold and Mikkelson 2002, Presley et al. 2010, Ulrich and Gotelli 2013), and identifies patterns on the sole basis of occurrences/absences. Ordination techniques such as correspondence analysis and multidimensional scaling naturally include species abundances and environmental data (Legendre and Legendre 1998). However, these techniques are not designed to infer the precise pattern of species associations (Dallas et al. 2016), and the basic definitions of turnover, modularity, and nestedness do not refer to environmental covariates. Consequently, nearly all studies on co-occurrence analysis focused on the presence–absence matrix only.

Co-occurrence is frequently linked either to species interactions in the case of modular and segregated patterns (Bascompte 2009, Götzenberger et al. 2012), or to environmental gradients (Ulrich et al. 2009) and colonisation trajectories (Gassert et al. 2013, Louy et al. 2014) in the case of nested patterns. We feel that this conceptual mismatch hinders progress in meta-community analysis, and we argue that there is need for an extension of the basic analysis of co-occurrences to incorporate environmental variables and possibly species functional traits. In this paper we develop such an extension to redefine species turnover, modularity and nestedness.

e second major shortcoming of binary matrix analysis lies in the plethora of metrics used to quantify co-occurrence patterns (Ulrich and Gotelli 2013). e most common metric of nestedness, NODF (Almeida-Neto et al. 2008), depends on specific ways of ordering the rows and columns of the matrix, whereas metrics of turnover, segregation, b-diversity (Stone and Roberts 1990, Tuomisto 2010), and modularity (Newman and Girvan 2004) are independent of matrix order. Because these metrics are based on dieferent algorithms and use di erent information from the matrix structure, results are not directly comparable. For instance, a recent meta-analysis (Ulrich and Gotelli 2013) confirmed that the degrees of turnover and nestedness are significantly negatively correlated. However, this correlation explained less than 20% of variance along the assumed nestedness – turnover gradient, suggesting the existence of additional intermediate patterns elusive to both metrics.

e third problem with the analysis of binary matrices is that raw scores of all these metrics heavily depend on matrix shape and fill and thus cannot be compared directly (Gotelli and Ulrich 2012). Consequently, pattern identifccation involves a second step, a comparison with an appropriate ecological standard, often in form of a randomised matrix (Gotelli and Graves 1996). Large  $e$  ect sizes reveal the existence of the focal pattern, while small e ect sizes imply a lack thereof. eoretically, it should be possible to define a gradient (based on e ect sizes) spanning, for instance, from strongly nested – to random – to strongly anti-nested. However, because there are three idealized patterns, it is problematic to place an assemblage on a single gradient from segregated to nested by using e ect sizes measured relative to a specifcc null model. Furthermore, to date, there is no procedure permitting to place a given meta-community unequivocally within the triangle defined by the modular, nested, and turnover vertices (Fig. 1).

ere have been several attempts to provide a unifying framework (Leibold and Mikkelson 2002, Presley et al. 2010) of meta-community structure. However, these attempts have not found a solution to the issue, nor they have included additional environmental information aimed at linking the observed patterns to ecological processes. We argue that, in order to make further progress in the classification of communities, we need to incorporate both environmental data, and species functional traits (McGill et al. 2006). Here we propose an expanded and unifying framework for the identifccation and measurement of nestedness, modularity, and turnover that surpasses the present methods of co-occurrence analysis by incorporating data on environmental characteristics and species traits. We show the simultaneous analysis of co-occurrences and environmental data provides new insight into the mechanisms of community assembly and the spatial distribution of species niches.

# **Methods**

# **A multidimensional unifying concept for the study of meta-community structure**

Eigenvector ellipses and their higher dimensional counterparts, ellipsoids, are commonly used in statistics (Friendly et al. 2013) and ecology (Legendre and Legendre 1998, Doledec et al. 2000) to visualise and quantify the spatial arrangement of sets of variables in two- and multi-dimensional space. Specifically, in principal components analysis (PCA), these ellipsoids are generated from the variance–covariance matrices of a set of equally weighed variables for which the associated eigenvectors and eigenvalues depict, respectively, the direction in space, and the semi-radii. In ecological analysis, these eigenvector ellipsoids were used to identify a functional niche space within an -dimensional hyperspace defined by a focal set of species characters (Hutchinson 1978, Jackson et al. 2011, Zalewski et al. 2014, Maire et al. 2015, Qiao et al. 2016). When applied to environmental measures collected at each site, our ellipsoid approach shows how species are distributed across multiple environmental dimensions. By identifying the region within this hyperspace occupied by each species, the ellipsoids can be interpreted as the realised range of environmental conditions within which a focal species occurs (Doledec et al. 2000, Janžekovi and Novak 2012). More notably, they provide novel ways to derive metrics of species turnover, modularity or nestedness. Building on this idea, here we use eigenvector ellipsoids (as defined in the Supplementary material Appendix 1) to redefine the concepts of nestedness, modularity, and turnover in terms of environmental characteristics or geographic positions of a set of sites hosting a meta-community. e corresponding Fortran software application 'NicheNew' is available from

W. Ulrich by request. We provide the respective source code in the Supplementary material Appendix 3.

### **Environmental based patterns of co-occurrence**

Our framework combines the  $\times$  presence–absence matrix ( species, sites) of species occurrences (**M**), and either a  $2 \times$  matrix of the geographical position of sites (typically longitude and latitude)  $(D)$ , or a  $\times$  matrix of

environmental variables (V) characterising each site. ese matrices are used to define and to compare species-specific eigenvector ellipsoids.

As shown in the worked example in Supplementary material Appendix 3 Fig. 3.1, for each species, the variables in **D** or **V** are normalised by the common Z-transformation to have a mean of zero and a standard deviation of one in order to ensure comparability between variables. en, for each species, a C matrix of dimensions  $\times$  is obtained by reducing **D** or **V** to a submatrix that includes only the sites of occurrence of the -th species (Fig. 1). Each **C** matrix is then used to define an ellipsoid  $E(C)$ , which identifies the realised environmental niche (the range of environmental conditions) (Fig. 1, 2) for each one of the species as:

$$
E(C) = \{ \epsilon R; (-c)^T \mathbf{E}^1(-c) \leq \epsilon^2 = (-c)^T U^{T-1} U(-c) \leq \epsilon^2
$$
 (1)

where <sup>2</sup> is the 99% a quantile of a  $\chi^2$ -distribution with degrees of freedom. e vector **c**, indicating the ellipsoid

centre, has coordinates  $(c_p,...,c)$  with  $c = \frac{1}{n} \sum_{i=1}^{n}$ 

 $= 1,...,$  Here  $n_j$  is the number of occurrences of the focal species, and  $v_i$  is the environmental value associated with each occurrence. e matrix  $\Sigma$  is the  $\times$  variance – covariance matrix of environmental variables (Supplementary material Appendix 1).

Equation 1 reveals whether or not a vector representing the environmental features (or the geographic coordinates) at a given location belongs to a given environmental niche. Specifically, our approach quantifies is es5187rall ee

niche overlap between pairs of species as exemplified in the supplementary material (Supplementary material Appendix 3 Fig. A3.1). e directions of axes of the -dimensional ellipsoid  $E(c)$  are given by the unit eigenvectors  $I_1, \ldots, I_n$  of the matrix  $\Sigma$ ; these vectors are the columns of the *U* matrix. e half-lengths  $\lambda = 1,..., \lambda$  of axes are defined by  $\lambda = \sqrt{\lambda^2 \lambda^2}$ . where  $\lambda$  is the eigenvalue of  $\Sigma$  and is contained in the diagonal  $\times$  matrix  $\Lambda$  (Supplementary material Appendix 1).

We can now redefine perfect nestedness as a pattern in which, for any pair of species, the eigenvector ellipse of one species is fully contained within that of the second species (Fig. 1). is behaviour results in a nested Russian-doll pattern when applied to a species rich meta-community. e common true subset definition of nestedness (Patterson and Atmar 1986) appears as a special case of this generalised definition.

Accordingly, we define perfect species turnover as a pattern in which the eigenvector ellipses of any pair of species do not overlap. Again, this is an extension of the common definition of turnover based on presence–absence data only (Fig. 1). Perfect turnover is the opposite of perfect nestedness, but can be achieved only if  $\tilde{\phantom{a}}$  . At higher matrix fill, the eigenvector ellipses necessarily partly overlap, and turnover is associated with some degree of species aggregation/modularity due to spatially separated groups of species having high ellipsoid overlap (O).

e degree of modularity can be quantified by the coe cient of variation  $CV = \frac{\sigma}{\mu}$ , where  $\mu$  is the average Euclidean distance between ellipsoid centroids within the hyperspace spanned by the Z-transformed environmental variables and  $\sigma$  its standard deviation. High values of  $CV_{env}$ are associated with increased modularity. As a reference, a Poisson random distribution is characterized by  $CV_{env}$  1.

is metric distinguishes turnover and modularity from the placement of the ellipsoid centres. e turnover pattern is characterised by equally distributed ellipses in space (Fig. 1) resulting in comparably low values of  $CV_{env}$ .

erefore, the ellipsoid approach makes it possible to identify three basic patterns in ecological presence–absence matrix. e results can be conveniently summarized in a plot of the coe cient of variation  $CV_{env}$  versus ellipse overlap  $O_{\text{env}}$  (Fig. 1). is representation provides an intuitive way to discriminate between nestedness (high  $O_{env}$  and low  $CV_{env}$ ), turnover (low  $O_{env}$  and low  $CV_{env}$ ), and modularity (intermediate  $O_{env}$  and high  $CV_{env}$ ) (Fig. 1).

#### **Functional trait based patterns of co-occurrence**

e ellipsoid approach also applies to species trait data (cf. Lamanna et al. 2014 for a related approach), coded in a  $\times$  species  $\times$  trait matrix **T**. As for the environmental variables the trait variables must allow the calculation of variance–covariance matrix, in this case based on interspecific variability. e transposes of the **M** and **T** matrices, with dimensions  $\times$  and  $\times$  , respectively, define the size of the trait ellipsoids of each site calculated over all occurring species, which constitutes a measure

Vegetation surveys, conducted according to Braun-Blanquet (1964) sampling procedures, took place from June to August 2003 to 2005. From these data, we compiled a single presence–absence data matrix  $M_{total}$  documenting the occurrence of 90 understorey plant species (matrix rows) across 130 plots (matrix columns; Supplementary material Appendix 2).

In each of the 130 studied plots, standard humus horizon soil samples were collected (Operat 2003, 2004, 2005). For the present study, we estimated soil fertility from the plot average of pH, N content, and the respective carbon– nitrogen (C/N) ratio. Species functional traits were assessed by Ellenberg indicator values (Ellenberg et al. 1992), with light, moisture, and soil nitrogen demands calculated by JUICE 7.0 (Tichý 2002) (Supplementary material Appendix 2). For simplicity we refer to them as functional traits although they rather reflect resource demands. Distances in trait space were assessed using the centroid distances of of the realised total trait space per s5(s)th15(r)-15(i)-15(a)-1<del>thlisedpeein&&lghMalle</del> elfipsbrds. Of course the results of our approach depend on trait choice. erefore, a careful choice of those traits that potentially influence the pattern of species co-occurrence is crucial for a proper interpretation of the results.

#### **Data analysis**

To relate environmental conditions to the pattern of species co-occurrence and environmental ellipsoid overlap  $(O_{\text{env}})$ , we used a sliding window approach, moving a window of eight columns along the 130 forest plots (with overlap of four columns). Prior to this operation, the plots were sorted according to the first eigenvector of the geographic distance matrix. is sorting ensured that the 31 resulting windows contained nearby plots with comparably similar environmental characteristics. Is procedure ensured su cient environmental variability among the windows to obtain reliable regression results, whereas a comparable bootstrap resampling approach would have combined plots regardless of their respective spatial and environmental distances. We used the same window approach for species traits and moved a sliding window (width eight, overlap four species) along the rows of species. is resulted in 15 windows combining species of similar numbers of occurrences.

Intuitively, raw scores of ellipsoid overlap and centroid distance should be a ected by matrix size and fill. To test for these e ects, we first applied a general linear model (with identity link function and equiprobable error structure) to assess (using all species pairs) the dependence of pairwise ellipsoid overlap and centroid distances on: the numbers of occurrences of the two species; fill of the matrix; and the respective pair-wise Sørensen similarity index. is was computed as 2  $/(\frac{1}{1} + \frac{1}{2})$ , where *is the number of* joint occurrences and  $\overline{1}$  and  $\overline{2}$ , the numbers of occurrences of both species. Additionally, we calculated for each sliding window the C-score as a measure of matrix-wide species segregation (Stone and Roberts 1990) and the NODF score (Almeida-Neto et al. 2008), which quantifies the degree of ordered species loss among plots (the degree of nestedness).

In species co-occurrence analysis, statistical inference is commonly based on a null model approach (Gotelli 2000, Gotelli and Ulrich 2012) to eliminate influences of matrix fill and size. We tested the performance of three well-established null models (Gotelli 2000, Gotelli and Ulrich 2012), and compared the above described metric scores with those obtained from 19 the fixed – fixed (FF) randomization based on the independent swap algorithm, 2) the equiprobable – equiprobable (EE) randomization based on an equiprobable random reshu ing of matrix occurrences, and 3) an equiprobable reshu ing (SH) of each environmental variable among columns. We choose these three algorithms because they span from very permissive (EE, which does not account for dieferential occurrence probabilities that are not linked to the focal pattern) to very conservative (FF and SH, which assume observed numbers of occurrences to be independent of the focal pattern).

Each analysis was based on 200 null matrices. We used raw scores of *O*env and *CV*env and respective null model e ect sizes (  $=Ob$  *E*; *Ob* and *E* : observed and expected scores) as response variables in a general linear

based on presence–absence data were assessed from standardized e ect sizes (SES =  $/\sigma$ ;  $\sigma$  is the standard deviation of the null model distribution) and similarly linked to soil variables.  $E$  ect sizes of predictors were estimated by partial  $\eta^2$ -values. Because the results obtained with FF and SH were qualitatively similar to those obtained with EE, we show the EE results only and provide the respective results from FF and SH in the electronic supplement D.

# **Results**

#### **Ellipsoid overlap and species co-occurrences**

Presence–absence based co-occurrence analysis revealed for the majority of shifting windows a significantly aggregated pattern of co-occurrence in comparison to the EE null model (average SES C-score  $=-12.9\pm0.53$ , average SES NODF =  $1.29 \pm 0.29$ ; mean  $\pm$  SE, Fig. 3c) and a segregated pattern in comparison to the FF null model (average SES C-score =  $4.16 \pm 0.62$ , average SES  $NODF = -2.34 \pm 0.71$ .

plants was partly determined by soil factors (Table 2, Supplementary material Appendix 4 Table A4.1). We found that nitrogen content and the respective variability among plots increased ellipsoid overlap, but did not influence distances among ellipsoid centroids (Table 2). In turn, pH and variability in pH were positively correlated with the distances among ellipsoid centroids. Neither of these patterns was detected by traditional species based co-occurrence analysis (Table 2). High C/N ratios and consequently low nitrogen availability decreased average distances among ellipsoid centroid (Table 2) and therefore the degree of species aggregation. Again, this pattern was not detected by the occurrencebased analysis. Compared to EE and FF, the SH null model proved to be conservative, and did not point to environmental influence of ellipsoid overlap and distances (Supplementary material Appendix 4 Table A4.2). High pH  $(r=0.54,$  $P(r=0) < 0.01$ ) and the high C/N ratio (r = 0.40, P(r = 0) 0.05pan.412ao7d/o412a.d/o4112a.d/o4112/o415H58 0 Td  $[0.041c$ Text

2007). Consequently, our results do not point to competition as a major driver of community organisation. We also have no evidence for high colonisation rates from a regional species pool and a resulting neutral community organisation. However, because community structure is known to change with spatial scale (Cottenie 2005), our interpretation does not preclude strong within-plot competitive interactions at a spatial scale below the resolution of our data.

Our approach allows for a straightforward and readily interpretable link between trait and environmental niche spaces. Previously, such links were provided by (mostly) linear ordination techniques like correspondence analysis or nonmetric multidimensional scaling. However, these methods did not di erentiate between the specific patterns of species co-occurrence. ey did also not directly refer to species functional niches. Because the degree of trait modularity was negatively correlated with the segregation of environmental ellipsoids (and hence niche space; Fig. 4c), we argue that environmental filtering sorted species according to environmental factors, which led to segregation in trait space. Such a pattern is expected if species with similar environmental demands have also similar functional traits. In this respect, we found a clear indication of modularity in the distribution of functional traits (Fig. 4a).

Interestingly, the degree of modularity was positively linked to average ellipsoid overlap. us, high overlap in environmental niche space, and therefore high species packing, might be associated with the appearance of species groups with similar functional traits. In turn, this might increase community stability, as reported by Bastolla et al. (2009) and Rohr et al. (2014) for mutualistic networks. However, specific studies aimed at linking trait distribution to species richness are lacking. In line with this finding, we observed decreased trait overlap with increasing environmental distance (Fig.  $4b-c$ ). Such a pattern is expected if environmental filtering also filters for traits resulting in a distance decay of trait overlap along environmental gradients (Kraft et al. 2015).

Here we focused on presence–absence data. However, our approach can be easily extended to incorporate abundance data by weighting the environmental variables by abundances in the second mapping step of analysis. Furthermore, our method can be applied to dataset including phylogenetic information. In fact, because the ellipsoid method is able to directly link phylogenetic distances to environmental niche overlap, it has also the potential to identify those environmental or geographical variables that contribute most to phylogenetic di erentiation.

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